

# Global Spillovers from FED Hikes and a Strong Dollar: The Risk Channel

José Cristi, Şebnem Kalemli-Özcan, Mariana Sans, and Filiz Unsal\*

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We study the international transmission of U.S. monetary policy (FED hikes) and a strong U.S. dollar. Both of these variables are endogenous and thus we follow the recent developments in the literature to measure the exogenous components of each from the perspective of the rest of the world (ROW). We follow the high frequency identification of [Gertler and Karadi \(2015\)](#) in order to capture the exogenous changes in the U.S. monetary policy around a short window of FOMC announcements by observing short-term asset prices in a 45-minute window. We follow [Obstfeld and Zhou \(2022\)](#) and [IMF ESR \(2023\)](#) and use the dollar’s appreciation against G10 currency as a global “dollar shock”.

We focus on the effects of these shocks on heterogeneous risk premia in emerging markets (EMs) vs advanced economies (AEs), measured by deviations from uncov-

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\*Forthcoming in AEA P&P, 2024. Cristi: University of Maryland, [jicristi@umd.edu](mailto:jicristi@umd.edu). Kalemli-Özcan: University of Maryland, [kalemli@umd.edu](mailto:kalemli@umd.edu). Sans: University of Maryland, [msans@umd.edu](mailto:msans@umd.edu). Unsal: OECD, [Filiz.UNSAL@oecd.org](mailto:Filiz.UNSAL@oecd.org). The views expressed and arguments employed in this paper are solely those of the authors and do not necessarily reflect the official views of the Organisation for Economic Co-operation and Development or its member countries. Any errors or omissions are the responsibility of the authors.

ered interest parity (UIP). As shown by [Kalemli-Özcan and Varela \(2021\)](#) and [Akinci, Kalemli-Özcan and Queralto \(2022\)](#), UIP premia is a country-level time-varying risk premia priced by the international investors to hold risky EM assets. [Albagli et al. \(2024\)](#) argue that a combination of idiosyncratic and global shocks lead international investors to price this risk premia differentially by country. Such time-varying deviations from the UIP have been identified as crucial in understanding deteriorating macro conditions in emerging markets with risk-sensitive capital flows ([Kalemli-Özcan \(2019\)](#), [Di Giovanni et al. \(2022\)](#), [Basu et al. \(2023\)](#), [Kalemli-Özcan and Unsal \(2023\)](#)). The UIP deviations can also be important for advanced economies as they are linked to demand for U.S. treasuries (e.g. [Degasperis, Hong and Ricco \(2023\)](#), [Jiang, Krishnamurthy and Lustig \(2021\)](#)).<sup>1</sup> To uncover the global impact of U.S. monetary policy and the dollar, we rely on local projections, as proposed by [Jordà \(2005\)](#). The local projection method provides a flexible framework and is easy to implement. Moreover, it is well documented that local projections have several advantages over VAR models (e.g. [Kilian and Lütkepohl \(2017\)](#); [Plagborg-Møller and Wolf \(2021\)](#)).

We find that, a tighter U.S. monetary policy leads to higher UIP premia in EMs, whereas such effects are not observed in AEs. Due to currency and credit risk, foreign investors demand a risk premium to absorb local currency debt compared to dollar debt. Hence another interpretation of the increase in the UIP premium is a temporary tightening of foreigners’ portfolio constraints in the absorption of local currency debt which may force private deleveraging and growth slowdown, compounded by fire sales in local currency assets (see [Basu et al. \(2023\)](#) for theoretical underpinnings of this channel, where closing the UIP wedge improves welfare). The U.S. monetary policy shock, which is financial in nature, brings an immediate depreciation along with a

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<sup>1</sup>See also quantitative models, where exogenous UIP deviations take center stage, such as [Dedola, Rivolta and Stracca \(2017\)](#), [Akinci and Queralto \(2023\)](#), [Gourinchas \(2018\)](#) for contractionary effects of the U.S. monetary policy on real outcomes of other countries.

higher interest rate differential between EMs and the U.S., as UIP risk premium is priced-in the interest rate differentials (as shown by [Kalemli-Özcan and Varela \(2021\)](#)). Thus, an expected appreciation is not what is driving the higher UIP premia—on the contrary, investors expect persistent depreciation after financial shocks on EM currencies as shown by [Kalemli-Özcan and Varela \(2021\)](#).

In contrast, in AEs, the UIP premia does not change, as expected appreciation and lower interest rate differentials cancel each other out. So the important difference between EMs and AEs is the opposite direction movement in the interest rate differentials as a response to exogenous FED hikes. The intuition for these results comes from [Kalemli-Özcan \(2019\)](#), who showed that short-term government bond spreads increasing in emerging markets and decreasing in advanced countries, as a result of tighter U.S. monetary policy. Such increase in interest rate differentials is not a result of tighter monetary policy in EMs, as shown by [De Leo, Gopinath and Kalemli-Özcan \(2022\)](#) since EMs own monetary policy is counter-cyclical.<sup>2</sup> Panels A and C of Figure 1 document these effects.

The effects of the dollar shock are strikingly different, as shown in Panels B and D of Figure 1. Now, UIP premium is decreasing in EMs and increasing in AEs, and the movements are much smaller compared to the U.S. monetary policy shock. Interest rate differentials are decreasing in both set of countries, but there is a persistent expected depreciation in EMs, indicating that the dollar shock picks up real/fundamental shocks instead of a financial shock as in the case of the U.S. monetary policy. These type of real/fundamental shocks require macroeconomic adjustment in the form of more depreciated EM currencies in the medium term rather than temporary downturns associated

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<sup>2</sup>[Degasperi, Hong and Ricco \(2023\)](#) show empirically how risk premia and commodity channels interact in the transmission of U.S. monetary policy. In particular, risk premium underlines the financial channel and is key to explain the real effects while the commodity channel can explain inflation dynamics. When U.S. monetary policy tightens, global demand contracts, implying higher risk premia, whereas with a direct shock to commodity prices, such as an appreciation of dollar, inflation can be higher.

with financial shocks. As expected, therefore, the dollar shock is not priced-in as risk premia in EMs, unlike the U.S. monetary policy shock. This distinction is key for policymakers because in EMs with financial frictions the salience of the UIP premium after the U.S. monetary policy shock may imply the need to use other tools such as FX interventions, as opposed to not intervening at all for the dollar shock.<sup>3</sup>

Overall, our analysis underscores a key difference between FED hikes and a strong dollar and the need for caution in interpreting the impact of these shocks especially on EMs. Unlike U.S. monetary policy shocks leading to higher risk premia in EMs, the dollar shock may not be financial and could be capturing a real shock which requires external adjustment and no policy intervention.

## 1 Data

We work with a quarterly unbalanced panel with 59 countries from 1990q1 to 2019q4; 42 are EMs. We use 12 month UIP deviations, measured in logs as follows:  $(i_t - i_t^{US}) - (s_{t+12}^e - s_t)$ , i.e. the difference between the log interest rate differentials and the gap between log expected and spot nominal exchange rate. We use the twelve-month treasury rates from Bloomberg, twelve-month expectations of the exchange rate from Consensus Economics and spot nominal exchange rates from IFS. Our analysis also includes real GDP from the World Economic Outlook and use the CPI from the IFS. We drop hard pegs and dual market exchange rate countries (Ilzetzki, Reinhart and Rogoff (2022) classifications 1 and 6). Thus, we always work with an unbalanced panel composed of managed and pure floats at the time of their inclusion.

We use two shocks. The first one is the Nominal Major Currencies U.S. dollar index

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<sup>3</sup>See Basu et al. (2023) for a conceptual framework that underlines the types of shocks that matter for the optimal policy response by the emerging market economies. The IMF's Integrated Policy Framework and its Institutional View on The Liberalization and Management of Capital Flows incorporate findings from this conceptual model.

from FRED, which we refer to as the dollar shock. It is a trade-weighted dollar index against a basket of G10 currencies.<sup>4</sup> We normalize it to a 10% depreciation shock as in [Obstfeld and Zhou \(2022\)](#).

The second is an exogenous measure for the U.S. monetary policy shocks. We rely on [Gertler and Karadi \(2015\)](#) monetary policy surprises. These surprises are obtained from high frequency changes in interest rates around central bank policy announcements. The key identifying assumption is that monetary policy is predetermined over the event window and, hence, not affected by financial market reaction. In particular, we use [Gertler and Karadi \(2015\)](#) averaged monthly weighted raw surprises in three-month Fed Fund Futures (FF4) to instrument the twelve-month U.S. treasury rate, which we obtain from Bloomberg.<sup>5</sup> We provide additional details about the data in the Appendix.

## 2 Empirical Analysis

We follow [Jordà \(2005\)](#) and [Stock and Watson \(2018\)](#) to compute the dynamic effects of these shocks on twelve-month UIP deviations (in logs) for country  $c$  in period  $t + h$ . We rely on two baseline specifications. Specification (1) captures the dynamic effects of U.S. monetary policy shocks, whereas specification (2) captures the dynamic effect of the dollar shock following [Obstfeld and Zhou \(2022\)](#). Controls ( $x_{c,t-i}$ ) include four lags of the dependent variable, GDP growth and inflation differentials with respect the U.S. These controls are consistent with the literature on the international transmission of U.S. monetary policy shocks, as they control for the trade and financial channels of international transmission.

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<sup>4</sup>The currencies included in the Index are the euro, Japanese yen, Canadian dollar, UK pound sterling, Swiss franc, Australian dollar, and Swedish krona.

<sup>5</sup>The monetary policy shocks from [Gertler and Karadi \(2015\)](#) pass comfortably the weak instrument tests, and hence they are relevant in capturing exogenous changes in the U.S. monetary policy. We can provide test results upon request.

$$y_{c,t+h} = \alpha_c + \beta_h \hat{i}_t^{US} + \sum_{i=1}^{i=4} \eta_i x_{c,t-i} + \varepsilon_{c,t+h} \quad (1)$$

$$y_{c,t+h} = \alpha_c + \beta_h Dollar_t + \sum_i^4 \eta_i x_{c,t-i} + \varepsilon_{c,t+h} \quad (2)$$

As the dependent variable ( $y_{c,t+h}$ ), we include UIP deviations and its components, the interest rate differential and the expected change in exchange rate, all in logs. We summarize results in Figures 2 and 3 below for the components and Figure 1 shows the response of the UIP premia.

As summarized in the introduction, we find that the two shocks yield opposite results on risk premia. The U.S. monetary shock, leads to an increase in UIP risk premia in EMs and a decrease in AEs. The dollar shock, however, does not lead to an increase in UIP premia in EMs but it does in AEs, despite a much smaller magnitude.

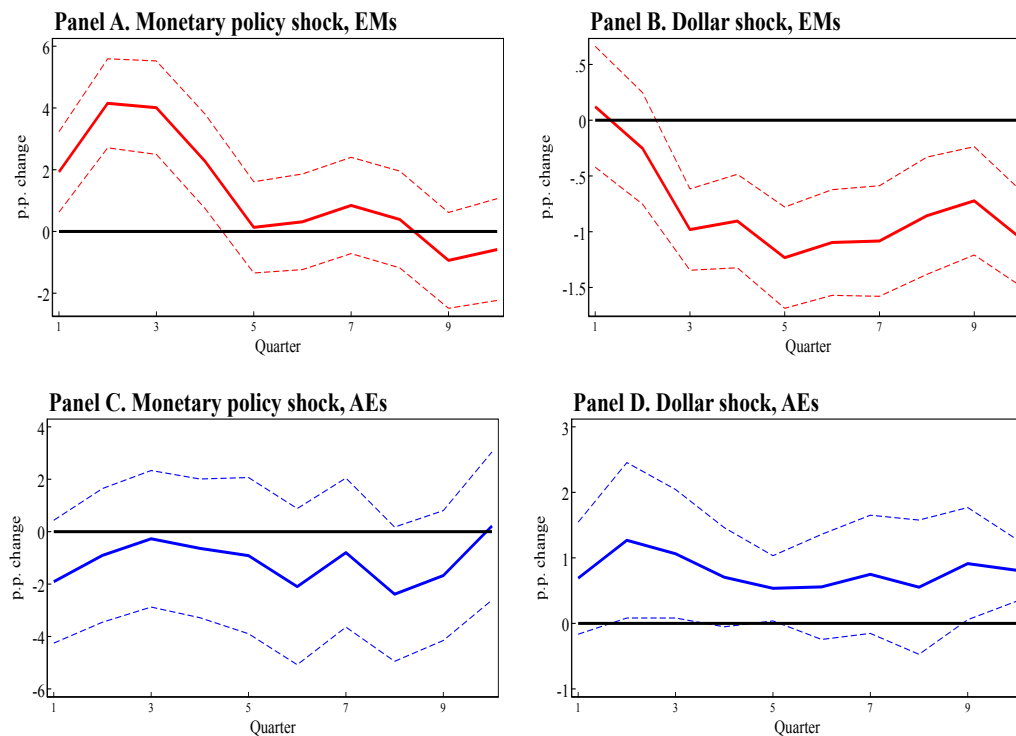
To understand why this is the case, we look at the components of UIP, namely the interest rate differential and the exchange rate adjustment. For the U.S. monetary policy shock (specification (1)), we find that for EMs, the interest rate differentials increase more than expected changes in exchange rate. In the case of the dollar shock, however, we show there is a persistent expected depreciation in spite of the actual depreciation at the time of the shock, indicating that the dollar appreciation captures more fundamental shocks which requires macroeconomic adjustment. Notably, the dollar shock does not create higher interest rate differentials, as this is not a financial shock and, hence, the risk premium does not change.

In AEs, both shocks work as expected, the exchange rate depreciates when the dollar appreciates (or when the U.S. hikes) but then there is an expected appreciation.

Results are robust when adding an additional set of global controls as in [Obstfeld and Zhou \(2022\)](#). In particular, in specification (1), we add the contemporaneous and

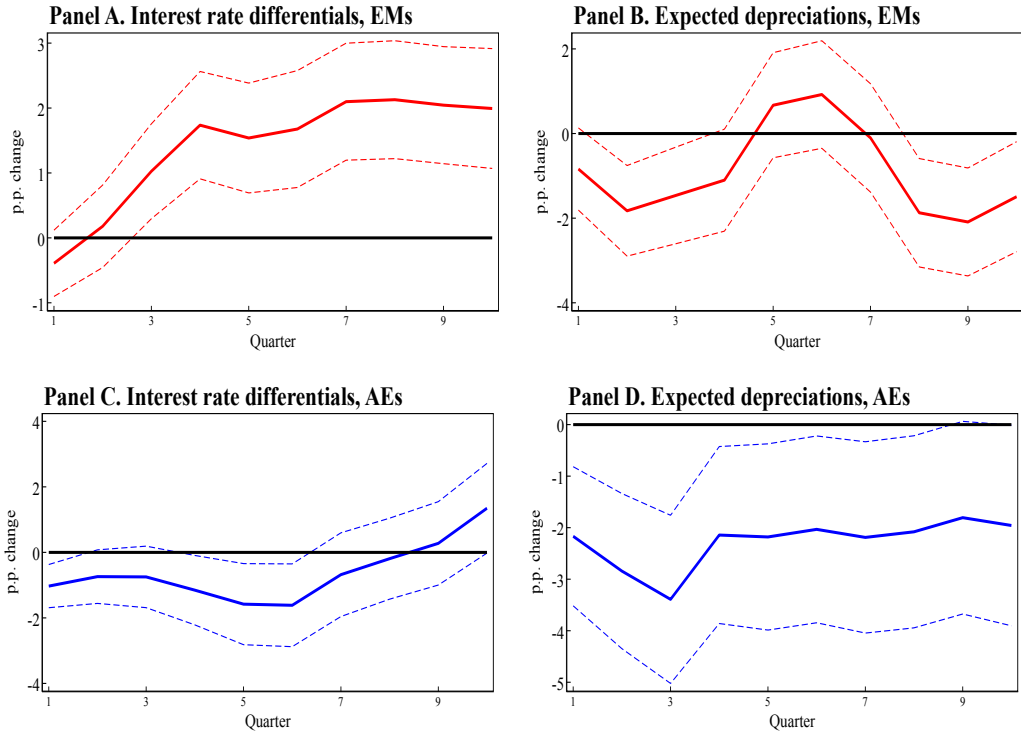
four lags of the oil price index, and median country trade balance, while in specification (2), we also include the twelve-month U.S. Treasury rate. We show these robustness results in the Appendix.

Figure 1: The UIP Premia: Response to U.S. Monetary Policy and Dollar Shocks



Notes: Figure shows the impulse response function of UIP premia to the monetary policy shock and the dollar shock, following specifications 1 and 2 respectively. Panels A and B are for EMs; panels C and D are for AEs.

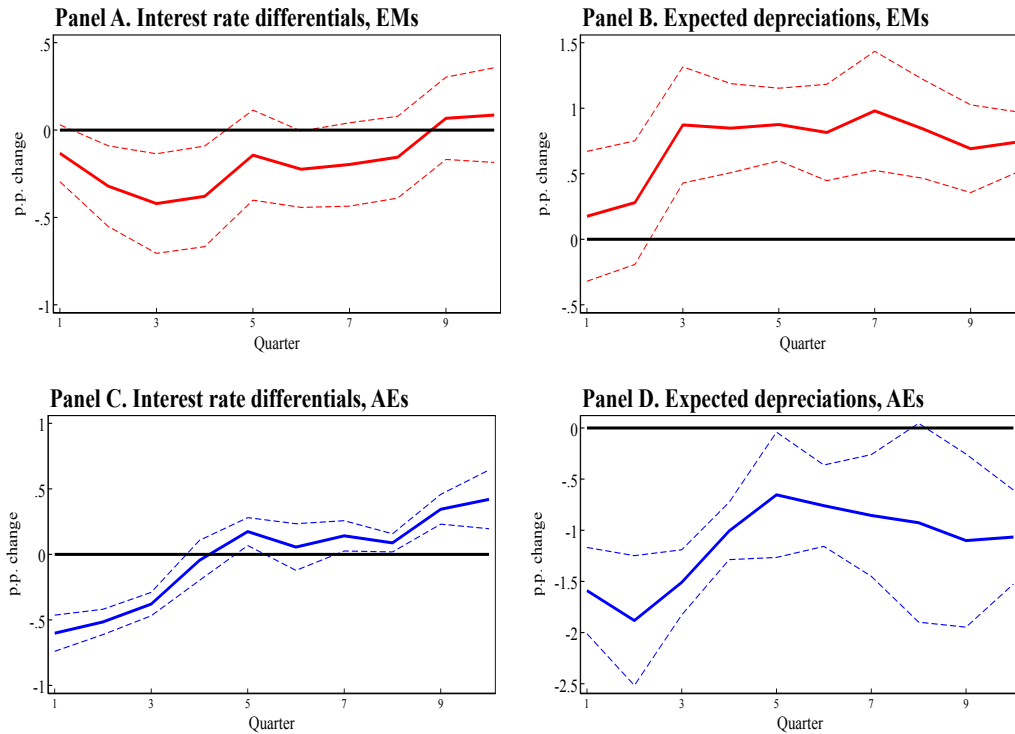
Figure 2: Components of the UIP Premia: Response to U.S. Monetary Policy Shocks



Notes: Figure shows the impulse response function of UIP components to the monetary policy shock, following specifications 1. Panels A and B are for EMs; panels C and D are for AEs.



Figure 3: Components of the UIP Premia: Response to Dollar Shocks



Notes: Figure shows the impulse response function of UIP components to the dollar shock, following specifications 2. Panels A and B are for EMs; panels C and D are for AEs.

### 3 Conclusion and Next Steps

We find that, unlike FED hikes, an appreciation of the U.S. dollar (the dollar shock) does not lead to higher risk premia in EMs, even though their currencies depreciate vis-a-vis the dollar. Our interpretation is that U.S. monetary policy shocks are directly linked to global financial conditions (e.g. [Rey \(2013\)](#)), whereas, a strong dollar may be capturing more fundamental, real (instead of financial) global shocks which requires external adjustment in EMs. Future work should look into underlying determinants for the differential effects of these shocks on the UIP risk premia, which is central to

GDP fluctuations in EMs under financial shocks.

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# A Appendix

## A.1 Data

Our data is of quarter frequency, and covers the period 1990q1-2019q4. In our analysis, we drop hard pegs and dual markets exchange rate countries, i.e. classifications 1 and 6 from [Ilzetki, Reinhart and Rogoff \(2022\)](#). We work with an unbalanced panel composed of managed and pure floats. We have a total of 59 countries in the sample which we use to run the EM vs AE exercises. We list the countries in [Table A1](#).

Below we describe the variables we use and we summarize data sources in [Table A2](#). Descriptive statistics are reported in [Table A3](#).

- 12m UIP deviation: calculated as the difference between log interest rate differentials and the gap between log expected and spot exchange rate, all at the same horizon. Log interest rate differentials are the short-term government bond or policy rate differentials vis-à-vis the United States. The log expected exchange rate is the 12-month ahead expected exchange rate as of month  $t$  and the log exchange rate is the spot rate (period average), both nominal and in terms of local currency per U.S. dollar.
- GDP: real seasonally adjusted
- CPI: period average
- Dollar shock: trade-weighted dollar index against a basket of G10 currencies from FRED (ticker DTWEXBGS). We use end of quarter observations and weights by merchandise trade weights.
- 12 month US treasury rate

- [Gertler and Karadi \(2015\)](#) shock: averaged monthly weighted raw surprises in 3-month Fed Fund Futures (FF4) from [Gertler and Karadi \(2015\)](#)

Table A1: Country Sample

Advanced Economies	Emerging Economies			
Denmark	Albania	Czech Republic	Mauritius	Slovak Republic
Finland	Argentina	Ecuador	Mexico	South Africa
Germany	Armenia	Egypt	Morocco	Thailand
Iceland	Azerbaijan	Guatemala	Pakistan	Tunisia
Ireland	Belarus	Hungary	Paraguay	Turkey
Israel	Brazil	India	Peru	Uruguay
Italy	Bulgaria	Indonesia	Philippines	
New Zealand	Chile	Kazakhstan	Poland	
Norway	China	Korea	Romania	
Spain	Colombia	Latvia	Russia	
Switzerland	Costa Rica	Malaysia	Serbia	
	Croatia	Malta	Singapore	

Notes: We follow the IMF 2000 World Economic Outlook country groups classification. Because we measure U.S. monetary policy spillovers, we drop the U.S.

Table A2: Data sources

Variable	Source
GDP	WEO
CPI	IFS
12m UIP deviation	Bloomberg, IFS and Consensus Forecast
US 12m treasury bill	Bloomberg
<a href="#">Gertler and Karadi (2015)</a> shock	Updated version of <a href="#">Gertler and Karadi (2015)</a>
Dollar shock	FRED

Table A3: Descriptive Statistics (1990q1-2019q4)

	mean	sd	min	max
GDP growth differential with US	0.004	0.024	-0.154	0.673
Inflation differential with US	0.016	0.030	-0.026	0.131
12m UIP deviation	0.028	0.041	-0.104	0.158
12m US treasury rate	0.032	0.023	0.001	0.083
GK(15) shock	-0.011	0.030	-0.179	0.056
Dollar shock	-0.005	0.334	-0.850	0.868

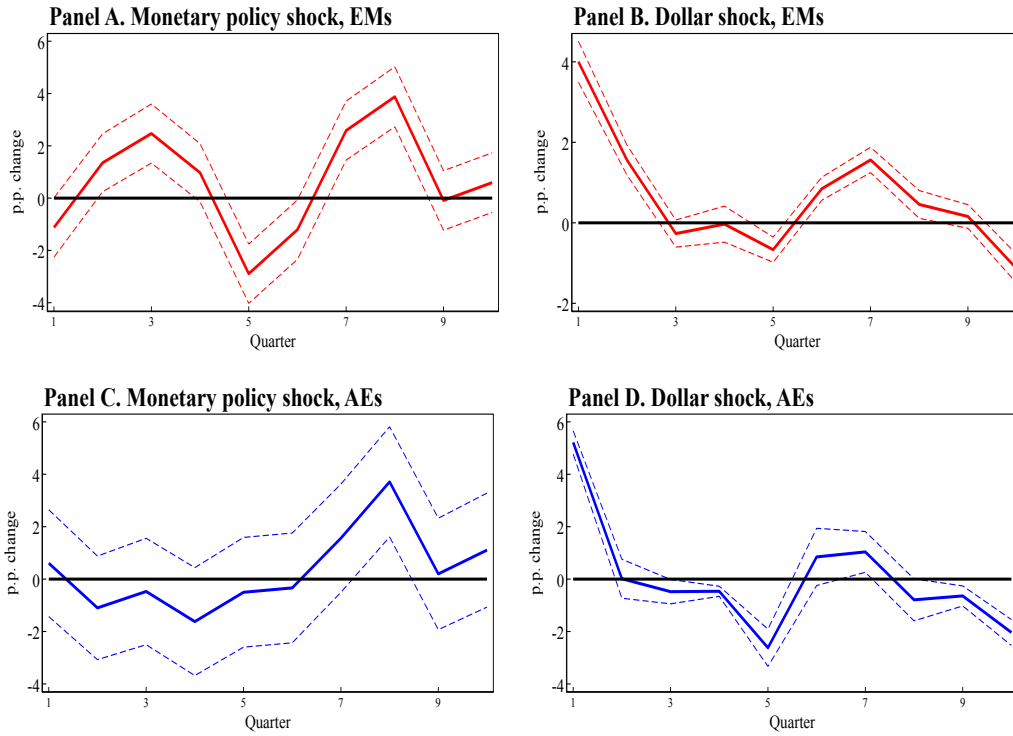
Notes: This table summarizes the descriptive statistics of the variables used in the empirical analysis for the period 1990q1-2019q4. Variables are as explained above.

## A.2 Additional Results

We also study the dynamics effects on nominal exchange rate. We show results in Figure [A1](#).



Figure A1: IRFs for Nominal Exchange rates for U.S. Monetary Policy and Dollar Shocks



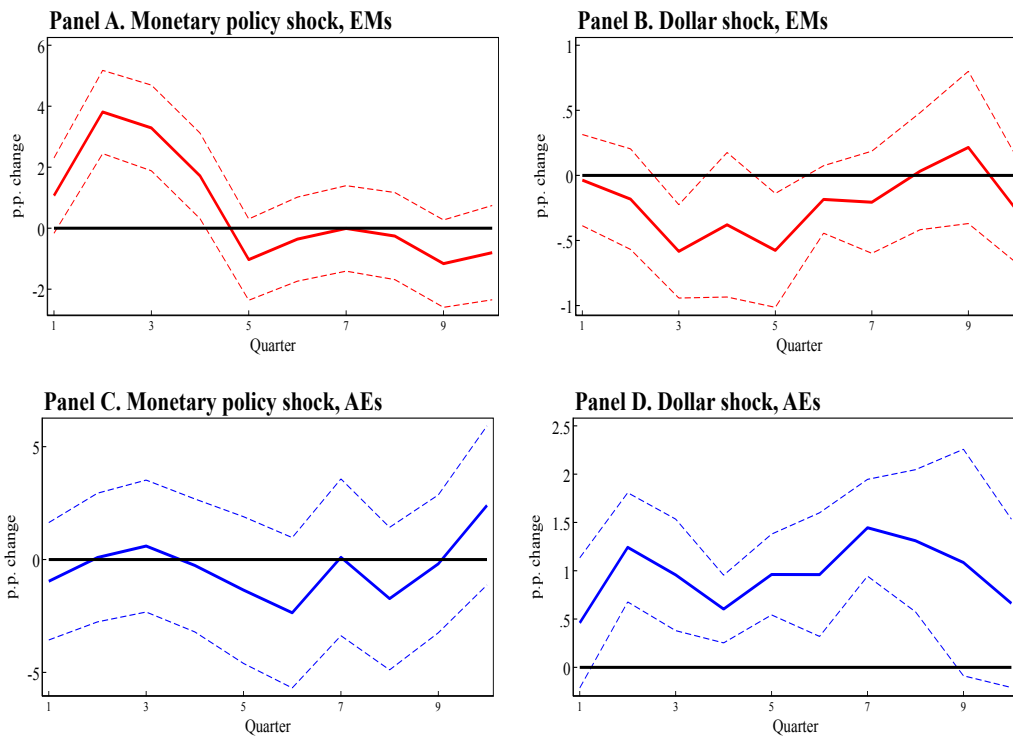
Notes: Figure shows the impulse response function of exchange rate to the monetary policy shock and the dollar shock, following specifications 1 and 2 respectively. Dependent variable is defined as the growth rate of nominal exchange rate (quarter to quarter). Panels A and B are for EMs; panels C and D are for AEs.

### A.3 Robustness

We include other global controls such as the oil price index from IMF, and the median of the trade balance within each group.<sup>6</sup>

<sup>6</sup>In particular we used POILAPSP index from IMF.

Figure A2: The UIP Premia: Response to U.S. Monetary Policy and Dollar Shocks w/Global Controls



Notes: Figure shows the impulse response function of UIP premia to the monetary policy shock and the dollar shock, following specifications 1 and 2. We include as controls the contemporaneous and four lags of: oil price index (logs) and the median trade balance within each group of countries. Panels A and B are for EMs; panels C and D are for AEs.