

# Lecture Notes on Canonical Correlation Analysis

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# Canonical Correlation Analysis

- Consider the random vector

$$X_{m \times 1} = \begin{pmatrix} U_{m_1 \times 1} \\ V_{m_2 \times 1} \end{pmatrix}$$

where

$$\begin{aligned} E[X] &= 0, \\ V(X) &= \Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \\ &\quad \begin{matrix} m_1 \times m_1 & m_1 \times m_2 \\ m_2 \times m_1 & m_2 \times m_2 \end{matrix} \end{aligned}$$

with  $\Sigma_{12} = \Sigma'_{21}$  and, without loss of generality, we assume that  $m_1 \leq m_2$ .

# Canonical Correlation Analysis

- Now, consider a linear combination

$$Y = \alpha' U$$

of the components of  $U$  and a linear combination

$$Z = \gamma' V$$

of the components of  $V$ .

- The objective is to find the linear combinations of  $U$  and  $V$  which give the maximum correlation.

# Canonical Correlation Analysis

- To proceed, note first that correlations are unique only up to a scalar multiplication since

$$\begin{aligned} |\text{Corr}(Y, Z)| &= \left| \frac{\text{Cov}(Y, Z)}{\sqrt{\text{Var}(Y)}\sqrt{\text{Var}(Z)}} \right| \\ &= \left| \frac{ab}{|ab|} \frac{\text{Cov}(Y, Z)}{\sqrt{\text{Var}(Y)}\sqrt{\text{Var}(Z)}} \right| \\ &= \left| \frac{\text{Cov}(aY, bZ)}{\sqrt{\text{Var}(aY)}\sqrt{\text{Var}(bZ)}} \right| \\ &= |\text{Corr}(aY, bZ)| \end{aligned}$$

for  $a, b \in \mathbb{R}$  such that  $a \neq 0$  and  $b \neq 0$ .

# Canonical Correlation Analysis

- It follows that, without loss of generality, we can normalize the variance of  $Y$  and  $Z$  as follows:

$$1 = \text{Var}(Y) = E[Y^2] = \alpha' E[UU'] \alpha = \alpha' \Sigma_{11} \alpha,$$

$$1 = \text{Var}(Z) = E[Z^2] = \gamma' E[VV'] \gamma = \gamma' \Sigma_{22} \gamma,$$

where  $\text{Var}(Y) = E[Y^2]$  and  $\text{Var}(Z) = E[Z^2]$  follow from the fact that  $E[Y] = \alpha' E[U] = 0$  and  $E[Z] = \gamma' E[V] = 0$ .

- Note that the above normalization is tantamount to setting

$$a = \frac{1}{\sqrt{\text{Var}(Y)}} \text{ and } b = \frac{1}{\sqrt{\text{Var}(Z)}}.$$

- Under this normalization, correlation between  $Y$  and  $Z$  is given by the covariance

$$E[YZ] = \alpha' E[UV] \gamma = \alpha' \Sigma_{12} \gamma.$$

# Canonical Correlation Analysis

- To maximize the correlation, we set up the Lagrangian

$$\mathcal{L}_1 = \alpha' \Sigma_{12} \gamma - \frac{1}{2} \lambda (\alpha' \Sigma_{11} \alpha - 1) - \frac{1}{2} \mu (\gamma' \Sigma_{22} \gamma - 1)$$

- The first-order conditions are

$$\frac{\partial \mathcal{L}_1}{\partial \alpha} = \Sigma_{12} \gamma - \lambda \Sigma_{11} \alpha = 0 \quad (1)$$

$$\frac{\partial \mathcal{L}_1}{\partial \gamma} = \Sigma_{21} \alpha - \mu \Sigma_{22} \gamma = 0 \quad (2)$$

- Premultiply (1) by  $\alpha'$  and (2) by  $\gamma'$ , we get

$$\alpha' \Sigma_{12} \gamma - \lambda \alpha' \Sigma_{11} \alpha = 0,$$

$$\gamma' \Sigma_{21} \alpha - \mu \gamma' \Sigma_{22} \gamma = 0.$$

# Canonical Correlation Analysis

- Next, differentiating  $\mathcal{L}_1$

$$\mathcal{L}_1 = \alpha' \Sigma_{12} \gamma - \frac{1}{2} \lambda (\alpha' \Sigma_{11} \alpha - 1) - \frac{1}{2} \mu (\gamma' \Sigma_{22} \gamma - 1)$$

with respect to  $\lambda$  and  $\mu$  and setting the partial derivatives equal to zero, we get

$$\begin{aligned} \frac{\partial \mathcal{L}_1}{\partial \lambda} &= -\frac{1}{2} (\alpha' \Sigma_{11} \alpha - 1) = 0 \text{ implying that } \alpha' \Sigma_{11} \alpha = 1 \\ \frac{\partial \mathcal{L}_1}{\partial \mu} &= -\frac{1}{2} (\gamma' \Sigma_{22} \gamma - 1) = 0 \text{ implying that } \gamma' \Sigma_{22} \gamma = 1 \end{aligned}$$

# Canonical Correlation Analysis

- It follows that

$$\begin{aligned}\alpha' \Sigma_{12} \gamma - \lambda \alpha' \Sigma_{11} \alpha &= \alpha' \Sigma_{12} \gamma - \lambda = 0, \\ \gamma' \Sigma_{21} \alpha - \mu \gamma' \Sigma_{22} \gamma &= \gamma' \Sigma_{21} \alpha - \mu = 0.\end{aligned}$$

Moreover, because  $\alpha' \Sigma_{12} \gamma = \gamma' \Sigma'_{12} \alpha = \gamma' \Sigma_{21} \alpha$ , we further deduce that

$$\lambda = \mu = \alpha' \Sigma_{12} \gamma.$$

- Putting everything together, we can rewrite the first-order conditions

$$\begin{aligned}\frac{\partial \mathcal{L}_1}{\partial \alpha} &= \Sigma_{12} \gamma - \lambda \Sigma_{11} \alpha = 0 \\ \frac{\partial \mathcal{L}_1}{\partial \gamma} &= \Sigma_{21} \alpha - \mu \Sigma_{22} \gamma = 0\end{aligned}$$

as

$$\begin{pmatrix} -\lambda \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda \Sigma_{22} \end{pmatrix} \begin{pmatrix} \alpha \\ \gamma \end{pmatrix} = 0 \quad (3)$$



# Canonical Correlation Analysis

- Now, for the system of equations

$$\begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} \begin{pmatrix} \alpha \\ \gamma \end{pmatrix} = 0$$

to have a nontrivial solution, i.e.,  $(\alpha', \gamma')' \neq 0$ , it must be that the matrix on the left-hand side is singular, so that

$$\det \begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} = 0 \quad (4)$$

- Since our objective is to maximize the correlation  $\lambda = \alpha'\Sigma_{12}\gamma$ , we choose  $\lambda_1$  to be the largest root of the determinantal equation above and let  $(\alpha'_1, \gamma'_1)'$  be the associated eigenvector.

# Canonical Correlation Analysis

- Next, we construct

$$Y_1 = \alpha'_1 U \text{ and } Z_1 = \gamma'_1 V$$

so that  $Y_1$  and  $Z_1$  are the normalized linear combinations of  $U$  and  $V$  which give the maximum correlation, with  $\lambda_1$  being the maximal (canonical) correlation.

- The analysis can then proceed further with finding a second linear combination of  $U$ , say  $Y_2 = \alpha'_2 U$ , and a second linear combination of  $V$ , say  $Z_2 = \gamma'_2 V$ , which have maximum correlation among all pairs not correlated with  $Y_1$  and  $Z_1$ .

# Canonical Correlation Analysis

- To proceed, we again set up a Lagrangian

$$\begin{aligned}\mathcal{L}_2 = & \alpha' \Sigma_{12} \gamma - \frac{1}{2} \lambda (\alpha' \Sigma_{11} \alpha - 1) - \frac{1}{2} \mu (\gamma' \Sigma_{22} \gamma - 1) \\ & + \nu \alpha' \Sigma_{11} \alpha_1 + \theta \gamma' \Sigma_{22} \gamma_1\end{aligned}$$

- In looking at the above Lagrangian, you may wonder why there are constraints only for  $\text{Cov}(Y_2, Y_1) = \alpha_2' \Sigma_{11} \alpha_1 = 0$  and  $\text{Cov}(Z_2, Z_1) = \gamma_2' \Sigma_{22} \gamma_1 = 0$  and not for either  $\text{Cov}(Y_2, Z_1) = \alpha_2' \Sigma_{12} \gamma_1 = 0$  and  $\text{Cov}(Z_2, Y_1) = \gamma_2' \Sigma_{21} \alpha_1 = 0$ . It turns out that if the former two constraints are satisfied, then the latter two are automatically satisfied.

# Canonical Correlation Analysis

- To see this, recall the first-order conditions from step 1 above

$$\begin{aligned}\frac{\partial \mathcal{L}_1}{\partial \alpha} &= \Sigma_{12}\gamma_1 - \lambda_1 \Sigma_{11}\alpha_1 = 0 \\ \frac{\partial \mathcal{L}_1}{\partial \gamma} &= \Sigma_{21}\alpha_1 - \lambda_1 \Sigma_{22}\gamma_1 = 0\end{aligned}$$

Given these relations, note that, if  $\text{Cov}(Y_2, Y_1) = \alpha_2' \Sigma_{11} \alpha_1 = 0$ , then

$$\text{Cov}(Y_2, Z_1) = \alpha_2' \Sigma_{12} \gamma_1 = \lambda_1 \alpha_2' \Sigma_{11} \alpha_1 = 0.$$

In addition, if  $\text{Cov}(Z_2, Z_1) = \gamma_2' \Sigma_{22} \gamma_1 = 0$ , then

$$\text{Cov}(Y_1, Z_2) = \gamma_2' \Sigma_{21} \alpha_1 = \lambda_1 \gamma_2' \Sigma_{22} \gamma_1 = 0.$$

# Canonical Correlation Analysis

- Next, note that the first-order conditions associated with  $\mathcal{L}_2$  are

$$\frac{\partial \mathcal{L}_2}{\partial \alpha} = \Sigma_{12}\gamma_2 - \lambda_2 \Sigma_{11}\alpha_2 + \nu_2 \Sigma_{11}\alpha_1 = 0 \quad (5)$$

$$\frac{\partial \mathcal{L}_2}{\partial \gamma} = \Sigma_{21}\alpha_2 - \mu_2 \Sigma_{22}\gamma_2 + \theta_2 \Sigma_{22}\gamma_1 = 0 \quad (6)$$

Premultiply (5) by  $\alpha_1'$  and (6) by  $\gamma_1'$ , we get

$$\begin{aligned} \alpha_1' \Sigma_{12} \gamma_2 - \lambda_2 \alpha_1' \Sigma_{11} \alpha_2 + \nu_2 \alpha_1' \Sigma_{11} \alpha_1 &= \nu_2 = 0, \\ \gamma_1' \Sigma_{21} \alpha_2 - \mu_2 \gamma_1' \Sigma_{22} \gamma_2 + \theta_2 \gamma_1' \Sigma_{22} \gamma_1 &= \theta_2 = 0. \end{aligned}$$

- Hence, the last two constraints in  $\mathcal{L}_2$  are not binding.

# Canonical Correlation Analysis

- It follows finding this second linear combination pair involves solving the same eigenvalue problem that was given by

$$\begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} \begin{pmatrix} \alpha \\ \gamma \end{pmatrix} = 0$$

and

$$\det \begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} = 0$$

In particular, let  $\lambda_2$  be the second largest root of the determinantal equation above and let  $(\alpha'_2, \gamma'_2)'$  be the associated eigenvector; then,

$$Y_2 = \alpha'_2 U,$$

$$Z_2 = \gamma'_2 V$$

are the normalized linear combination pair which yields the highest correlation among all pairs not correlated with  $Y_1$  and  $Z_1$  and  $\lambda_2$  gives the value of their correlation.

# Canonical Correlation Analysis

- We can proceed in this way to obtain

Canonical Correlations (Eigenvalues)	Weights (Eigenvalues)	$Y$	$Z$
$\lambda_1$	$(\alpha'_1, \gamma'_1)'$	$Y_1 = \alpha'_1 U$	$Z_1 = \gamma'_1 V$
$\lambda_2$	$(\alpha'_2, \gamma'_2)'$	$Y_2 = \alpha'_2 U$	$Z_2 = \gamma'_2 V$
$\vdots$	$\vdots$	$\vdots$	$\vdots$
$\lambda_{m_1}$	$(\alpha'_{m_1}, \gamma'_{m_1})'$	$Y_{m_1} = \alpha'_{m_1} U$	$Z_{m_1} = \gamma'_{m_1} V$

# Canonical Correlation Analysis

- As a final remark, we note that the equations

$$\begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} \begin{pmatrix} \alpha \\ \gamma \end{pmatrix} = 0$$

and

$$\det \begin{pmatrix} -\lambda\Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & -\lambda\Sigma_{22} \end{pmatrix} = 0$$

can be written in a different form which may be more convenient for cointegration analysis.



# Canonical Correlation Analysis

- In particular, note that, since  $\lambda = \mu$ , we can rewrite the first-order conditions given by equations (1) and (2) as

$$\Sigma_{12}\gamma - \lambda\Sigma_{11}\alpha = 0 \quad (7)$$

$$\Sigma_{21}\alpha - \lambda\Sigma_{22}\gamma = 0 \quad (8)$$

- Premultiplying expression (8) above by  $\Sigma_{12}\Sigma_{22}^{-1}$  yields

$$\Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}\alpha - \lambda\Sigma_{12}\Sigma_{22}^{-1}\Sigma_{22}\gamma = \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}\alpha - \lambda\Sigma_{12}\gamma = 0$$

or

$$\Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}\alpha = \lambda\Sigma_{12}\gamma = \lambda^2\Sigma_{11}\alpha$$

where the last equality above follows from (7).

# Canonical Correlation Analysis

- It follows that, alternatively, we can obtain the squared canonical correlations and the eigenvectors  $\alpha_i$  ( $i = 1, \dots, m_1$ ) as solutions of the system of equations

$$(\lambda^2 \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}) \alpha = 0$$

where  $\lambda_i^2$  ( $i = 1, \dots, m_1$ ), i.e., the squared canonical correlations, are the roots of the determinantal equation

$$\det(\lambda^2 \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}) = 0.$$

- Similarly, it can be shown that  $\lambda_i^2$  and  $\gamma_i$  ( $i = 1, \dots, m_1$ ) can be obtained as the solutions of the system of equations

$$(\lambda^2 \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}) \gamma = 0$$

with  $\lambda_i^2$  ( $i = 1, \dots, m_1$ ) being the  $m_1$  largest roots of

$$\det(\lambda^2 \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}) = 0.$$